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Towards Explainable and Fair Supervised Learning

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Abstract

Algorithms that aid human decision-making may inadvertently discriminate against certain protected groups. We formalize direct discrimination as a direct causal effect of the protected attributes on the decisions, while induced indirect discrimination as a change in the influence of non-protected features associated with the protected attributes. The measurements of average treatment effect (ATE) and SHapley Additive exPlanations (SHAP) reveal that state-ofthe-art fair learning methods can inadvertently induce indirect discrimination in synthetic and real-world datasets. To inhibit discrimination in algorithmic systems, we propose to nullify the influence of the protected attribute on the output of the system, while preserving the influence of remaining features. To achieve this objective, we introduce a risk minimization method which optimizes for the proposed fairness objective. We show that the method leverages model accuracy and disparity measures.

1. Introduction

Discrimination consists of treating somebody unfavorably because of their membership to a particular group, characterized by a *protected attribute*, such as race or gender. To prevent *disparate treatment*, the law often forbids the use of certain protected attributes, such as race or gender, Z, in decision-making, e.g., in hiring, and dictates that these decisions, Y, shall be based on relevant attributes, X, and not depend on the protected attribute, Z. Historically, e.g., in the case of *redlining*, the prohibition of such direct discrimination was sometimes circumvented by the use of attributes correlated with the protected attribute as proxies. This is a particularly acute problem for machine learning data-rich systems, since they often find surprisingly accurate surrogates for protected attributes when a large set of legitimate-looking features is available, resulting in the *inducement* of discrimination via association (Wachter, 2019). To prevent such inducements of discrimination, legal systems establish that the impact of a decision-making process should be the same across groups differing in protected attributes, unless relevant attributes justify it, according to a *business necessity clause* (BNC) (Title VII of the Civil Rights Act, 1964). The main challenge in introducing non-discriminatory learning algorithms lies in preventing the inducement of indirect discrimination, while simultaneously avoiding direct discrimination (Zafar et al., 2015).

Related works. In machine learning, discrimination is typically defined based on statistical independence or causal relations. Well-known fairness objectives, such as parity of impact and equalized odds, correspond to the statistical independence between Z and Y (Hardt et al., 2016; Zafar et al., 2017; Aswani & Olfat, 2019). However, these notions are inconsistent with their legal counterparts (Lipton & Steinhardt, 2019) as legal systems allow for crucial exceptions from this independence through the BNC which permits decisions, Y, to depend on Z through X.

Causal approaches define direct and indirect discrimination as direct and indirect causal influence of Z on Y, respectively (Zhang et al., 2017; Zhang & Bareinboim, 2018; Marx et al., 2019). While this notion of direct discrimination is consistent with the concept of disparate treatment in legal systems, the corresponding indirect discrimination is not consistent with them, since the BNC allows for the use of an attribute that depends on the protected feature (causally or otherwise). This issue is addressed by pathspecific notions of causal fairness (Nabi et al., 2019; Chiappa, 2019; Wu et al., 2019). These methods allow for fair causal paths in which the impact of the protected attribute is permitted, thus, allowing for BNCs. However, if there is no limit on the influence that can pass through such a path. then the path can be used for indirect discrimination, as in the aforementioned case of redlining.

Problem summary. Consider a model supporting human decisions trained on a dataset of n samples $D = \{(\mathbf{x}^i, \mathbf{z}^i, y^i)\}$, where $\mathbf{x}^i \in \mathcal{X}, \mathbf{z}^i \in \mathcal{Z}, y^i \in \mathcal{Y}$, and i = 1, ..., n. The goal of a standard supervised learning algorithm is to obtain a function $\hat{y} : \mathcal{X} \to \mathcal{Y}$ that op-

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1055 timizes a given objective, e.g., $\mathbb{E}[\ell(Y, \hat{y}(X))]$, where the 1056 expectation is over the samples in D and ℓ is a loss func-1057 tion. If the dataset is tainted by discrimination, the crucial 1058 question is how to drop Z from a model without inducing 1059 discrimination, that is without increasing the impact of rel-1060 evant attributes X correlated with Z in an unjustified and 1051 discriminatory way as in radiining

061 discriminatory way as in redlining.

062 Contributions. Our work bridges statistical and causal no-063 tions of fairness with the literature on explainability, while 064 staying consistent with legal systems. First we define the 065 concepts of direct and induced indirect discrimination via 066 measures of causal influence. Second, we construct loss 067 functions, grounded in causality and explainability litera-068 ture, that measure the change in influence of X while the 069 protected attribute Z is removed. Third, we introduce and 070 evaluate an optimization method that drops the protected Z from a model while minimizing the induction of indirect discrimination through the non-protected features Xby minimizing the aforementioned loss functions. 074

2. Problem formulation

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077 Consider decisions Y that are outcomes of a model that acts 078 on random variables W having support in W. The dimen-079 sions of W are indexed, e.g., W_i corresponds to the *i*'th random variable with support \mathcal{W}_i , where $i \in \mathcal{F}$. We distin-081 guish between a set of non-protected attributes \mathcal{N} , consti-082 tuting the $|\mathcal{N}|$ -dimensional random variable $W_{\mathcal{N}} = X$, 083 and a set of protected features \mathcal{P} , constituting the $|\mathcal{P}|$ -084 dimensional $W_{\mathcal{P}} = Z$. These sets are non-empty, non-085 overlapping, and the set of all features is $\mathcal{F} = \mathcal{N} \cup \mathcal{P}$. 086

087The model generating decisions Y can suffer the effects of088training on discriminatory data. We propose that a non-089discriminatory model, \hat{Y} , of Y shall remove the influence090of the protected features on Y, while preserving the influ-091ence of remaining features on Y. In the following sub-092sections, we develop loss functions for supervised learning093that aim to achieve this objective.

2.1. Formulation based on causal effect measures

Formal frameworks for causal models include classic potential outcomes (PO) and structural causal models (SCM) (Pearl, 2009) or more recent segregated graphs that include undirected causal relationships (Shpitser, 2015).
The methods presented in this work do not rely on the notion of intervention, which tends to have a consistent meaning across causal frameworks.

104 Note that decisions Y are causal outcomes of the model 105 and the causal parents of these decisions are W. This 106 crucial point, emphasized in causal explainability litera-107 ture (Janzing et al., 2019), allows us to compute influence 108 measures via causal interventions on chosen components of W, as if there was no direct causal links between the components of W. Following SCM framework, samples of Y are generated by some function, $y = f(w, \epsilon)$, where ϵ is exogenous noise. Since the exogenous noise is unpredictable, here we focus on the de-noised function $y(w) = \mathbb{E}_{\epsilon} f(w, \epsilon)$. In the notation of SCM and PO, the potential outcome for variable Y after intervention do(Z = z) is written as Y_z . Average treatment effect of z on y w.r.t. a reference intervention z' is defined via respective interventions (Pearl et al., 2016),

$$ATE_Y(\boldsymbol{z}', \boldsymbol{z}) = \mathbb{E}[Y_{\boldsymbol{z}'} - Y_{\boldsymbol{z}}] = \mathbb{E}[Y|\boldsymbol{X}, \boldsymbol{z}'] - \mathbb{E}[Y|\boldsymbol{X}, \boldsymbol{z}],$$

where the last two expectations are over ϵ and a marginal distribution of $P(\mathbf{X})$, due to the *causal adjustment* for \mathbf{X} . A non-causal estimate would use conditional $P(\mathbf{X}|\mathbf{z})$ instead of $P(\mathbf{X})$. The causal *controlled direct effect* of \mathbf{z} on y w.r.t. a reference intervention \mathbf{z}' and intervention \mathbf{x} is

$$CDE_Y(\boldsymbol{z}', \boldsymbol{z} | \boldsymbol{x}) = \mathbb{E}[Y_{\boldsymbol{x}, \boldsymbol{z}'} - Y_{\boldsymbol{x}, \boldsymbol{z}}]. \tag{1}$$

Definition 1. Direct discrimination is the causal influence of a protected attribute \mathbf{Z} on the decisions Y in the sense that $\exists_{z,z'\in\mathcal{P}} \exists_{x\in\mathcal{N}} CDE_Y(z, z'|x) \neq 0.$

To remove this discrimination, one can construct a model \hat{Y} that does not use Z. However, this may introduce indirect discrimination into the model via the non-protected attributes X_i associated with the protected attributes Z.

Definition 2. Indirect discrimination induced via X_i is a change in the influence of X_i that depends on Zbetween the causal process Y and its model \hat{Y} , i.e., $\exists_{z \in \mathcal{P}} \exists_{x,x' \in \mathcal{N}} CDE_Y(x, x'|z) \neq CDE_{\hat{Y}}(x, x'|z)$ such that $P(x|z) \neq P(x)$ or $P(x'|z) \neq P(x')$.

To preserve influence of non-protected attributes we can minimize the following loss

$$\begin{split} L_{\text{ATE}}^{\text{IND}}(\boldsymbol{X}) &= \sum_{i} L_{\text{ATE}}(X_{i}) = \\ \sum_{i} \mathop{\mathbb{E}}_{X_{i}^{\prime\prime},X_{i}} \ell(\text{ATE}_{Y}(X_{i},X_{i}^{\prime\prime}),\text{ATE}_{\hat{Y}}(X_{i},X_{i}^{\prime\prime})). \end{split}$$

A similar loss could be constructed based on the comparison between $\text{CDE}_Y(X, X''|Z)$ and $\text{CDE}_{\hat{Y}}(X, X''|Z)$. In this paper we focus on losses that compare ATE and SHAP input influence measures.

2.2. Formulation based on input influence measures

Alternatively, influence can be measured on the grounds of input influence measures introduced to explain black-box AI models.

To measure the influence of a certain variable W_i prior works suggest to make a probabilistic intervention on that

variable by replacing it with some W'_{4} (Datta et al., 2016; 111 Lundberg & Lee, 2017; Janzing et al., 2019). Let the random input variables be W = XZ, which are a concatenation of variables X and Z. Let primed variables have the same joint distribution as the non-primed variables, $\forall_{w \in W} P(W' = w) = P(W = w)$, while being independent from them, $W' \perp W$. Let double primed variables have the same marginal distributions as the non-primed variables, $\forall_{i \in \mathcal{F}} \forall_{w \in \mathcal{W}_i} P(W_i'' = w) = P(W_i = w)$, and be independent from each other and the non-primed variables, i.e., $\forall_{i \in \mathcal{F}} \forall_{j \neq i} W_i \perp W_j$, $W'' \perp W'$ and $W'' \perp$ W. Then, the random variable $W_T W'_{-T} = W_T W'_{\mathcal{F} \setminus T}$ represents a modified random variable W with its components W_i replaced with samples from $P(\mathbf{W}')$ for each $i \in \mathcal{F} \setminus T.$

For any subset of features T that does not contain i, we can define a marginal influence (Datta et al., 2016; Janzing et al., 2019)

$$\mathbf{MI}_{Y}(W_{i}|\boldsymbol{w},T) = \mathbb{E}_{\boldsymbol{W}'}\left[Y_{\boldsymbol{w}_{T\cup\{i\}}\boldsymbol{W}_{-(T\cup\{i\})}} - Y_{\boldsymbol{w}_{T}\boldsymbol{W}_{-T}}\right],$$

where W' is a random baseline.

A popular measure of input influence is based on the Shapley value, which averages the marginal influence over all possible subsets T (Datta et al., 2016; Lundberg & Lee, 2017),

$$\operatorname{SHAP}_{Y}(w_{i}|\boldsymbol{w}) = \sum_{T \subseteq \mathcal{F} \setminus \{i\}} \frac{\operatorname{MI}_{Y}(W_{i}|\boldsymbol{w}, T)}{|\mathcal{F}|\binom{|\mathcal{F}|-1}{|T|}}.$$
 (2)

To preserve influence of non-protected attributes we can minimize the following loss,

$$L_{\text{SHAP}}^{\text{IND}}(\boldsymbol{X}) = \sum_{i} L_{\text{SHAP}}(X_{i}) = \sum_{i} \mathcal{L}_{\text{SHAP}}(X_{i}|\boldsymbol{X}\boldsymbol{Z}''), \underset{\boldsymbol{Z}''}{\mathbb{E}} \operatorname{SHAP}_{\hat{Y}}(X_{i}|\boldsymbol{X}\boldsymbol{Z}'')).$$

3. Minimizing $L_{ATE}(X)$ and $L_{SHAP}(X)$

We seek models \hat{Y} of binary Y that remove the influence of the protected attributes Z, while preserving the influence of non-protected attributes X by minimizing $L_{ATE}^{IND}(X)$ or $L_{SHAP}^{IND}(X)$ via transfer learning. First, we drop the protected attribute(s) Z from the data. We then obtain "Trad. w/o Z" model by minimizing the cross entropy loss, $H(\hat{y}, y) = -\sum_i y_i \log \hat{y}_i$. Next, we optimize for either $L_{ATE}^{IND}(X)$ or $L_{SHAP}^{IND}(X)$. For both objectives we use ℓ_2 loss. We refer to these two-stage optimization-based methods as OPT-ATE and OPT-SHAP, respectively. The training is done using momentum based gradient optimizer ADAM (Kingma & Ba, 2017) via batch gradient descent. We fine-tune two hyper-parameters: learning rate (α) and number of epochs (N). During fine-tuning we pick the values for which we get the best performance on the validation set. In our datasets, α is 1e-3 to 1e-2 and N is from 20 to 100. Our implementations of the methods will be released publicly as a Python library.

4. Experiments

We examine our method's and other supervised learning methods addressing discrimination's performance in binary classification on synthetic and real-world datasets. We measure $\mathbb{E}_{\boldsymbol{X},\boldsymbol{Z}} |\text{SHAP}_Y(X_i|\boldsymbol{X},\boldsymbol{Z})|$, following Lundberg & Lee (2017), and $\mathbb{E}_{X_i,X_i'} |\text{ATE}_Y(X_i,X_i')|$. To reduce computational costs, we use sub-sampling to compute these. In addition, we measure accuracy, demographic disparity ($|P(\hat{y} = 1|z = 0) - P(\hat{y} = 1|z = 1)|$), and equal opportunity difference ($|P(\hat{y} = 1|y = 1, z = 0) - P(\hat{y} = 1|y = 1, z = 0)|$). The dataset is partitioned into 20:80 test and train sets. All results are computed on the test set.

4.1. Evaluated learning methods

We evaluate four learning methods addressing discrimination at different stages of a machine learning pipeline (abbreviations in parenthesis). **Pre-processing:** Reweighing approach from Kamiran & Calders (2012). **In-processing:** (1) Reductions model ("Exp Grad") from Agarwal et al. (2018). We evaluate four variations of reductions constraining demographic parity, equalized odds, equal opportunity, and error ratio ("DP", "EO", "TPR", and "ER"). (2) Adversarial debiasing ("Adv Deb") from Zhang et al. (2018). **Post-processing:** Calibrated equalized odds approach ("CalEqOdd") from Pleiss et al. (2017).

We use the implementations of these algorithms as provided in the AI Fairness 360 open-source library (Bellamy et al., 2018). The baseline "traditional" model and underlying classifier for all the evaluated models is logistic regression. We also evaluate a logistic regression model that drops the protected attribute, Z, before training.

4.2. Synthetic results

To generate the synthetic dataset we draw samples from a multivariate normal distribution with standard normal marginals and given correlations. We then convert a column of our matrix into binary values, set that as Z, and set the rest as X. The correlations between both (X_1, X_2) and (X_2, Z) are zero. We compare the learning methods while increasing the correlation $r(X_1, Z)$ from 0 to 1. We use a simple model, $Y = \sigma(X_1 + X_2 + Z + 1)$ where σ is the logistic function.

Both OPT approaches preserve X_1 's influence with respect to the full model as $r(X_1, Z)$ increases (red and solid blue lines in Figure 1). As expected, the influence of X_1 in-

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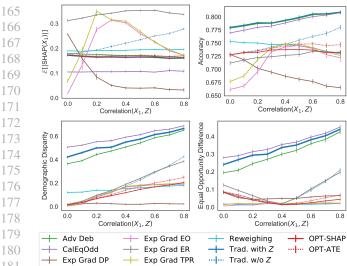


Figure 1. SHAP influence of X_1 , model accuracy, and two popular fairness measures as we increase the correlation $r(X_1, Z)$. Error bars show 95% confidence intervals based on 30 samples.

creases with correlation for the traditional method that simply drops Z, i.e., it induces indirect discrimination via X_1 (dotted blue line in Figure 1). Interestingly, even though the OPT does not optimize for either fairness measure, it performs better for all fairness measures than the traditional method dropping Z (in Appendix A we show results for two other fairness measures).

195 Other methods addressing discrimination either change the 196 influence of X_1 with the growing correlation $r(X_1, Z)$ 197 ("Exp Grad" in Figure 1) or use the protected attribute 198 Z and thus discriminate directly (see "Adv Deb", "CalE-199 qOdd", "Reweighing" in Appendix A). For instance, the 200 method optimizing for parity of impact decreases the im-201 pact of X_1 , because it aims to remove the correlation be-202 tween \hat{Y} and Z (brown line in Figure 1). Results for ATE 203 are qualitatively the same as for SHAP (Appendix C). 204

4.3. Real-world results

We train and test the evaluated methods on the COMPAS criminal recidivism dataset (Larson et al., 2016). Here, the 208 model predicts the recidivism of an individual based on 209 210 their demographics and criminal history with race being the protected attribute. To make the presentation more clear, 211 we exacerbate the racial bias by removing 500 samples of 212 positive outcomes (no recidivism) for African-Americans. 213 214 Data functions from the AIF360 library are used for this 215 dataset. Results for the unmodified German Credit dataset are qualitatively equivalent (see Appendix B). 216

In line with the synthetic results, the OPT approaches are not influenced by the protected attribute Z and, with respect

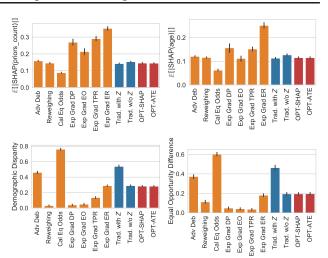


Figure 2. Averaged absolute SHAP for the two features most correlated with the protected attribute and fairness measures for the evaluated models on the COMPAS dataset. Error bars show 95% confidence intervals.

to the traditional model, preserve the influence for the two attributes most correlated with Z in this real-world scenario (blue and red in the top row of Figure 2). While most of the evaluated models outperform the OPT models for the fairness measures, they are either influenced by the protected attribute or do not preserve the influence of at least one of the most correlated attributes and have significantly lower accuracy (Appendix A). Therefore, as with the synthetic results, the changes in influence for these attributes indicate that these methods induce indirect discrimination during training, despite having better performance for certain fairness measures.

5. Conclusions

The presented results shed a new light on the problem of discrimination prevention in supervised learning. First, we propose a formal definition of induced discrimination, inspired by research in humanist fields (Altman, 2016) and discrimination via association (Wachter, 2019). We measure influence of features to capture induced discrimination. Second, we show that state-of-the-art methods addressing discrimination can return biased models influenced by the protected attribute or attributes associated with it when they are trained on potentially discriminatory datasets. Third, we propose an optimization-based method for discrimination prevention. The method drops the protected attribute and preserves the influence of nonprotected attributes to prevent the induction of discrimination via association. These results provide support for the use of the optimization approach in the circumstances where discrimination could have affected the training dataset.

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